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# Is cranial multi-detector computed tomography imaging valuable for stature estimation in Egyptian population?

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## Abstract

**Background and objectives:** Stature estimation is an initial and essential component of any medico-legal investigations. However, it becomes more challenging when only skull remains are available. So, the goal of this study is to assess stature estimation using cranial multi-detector computed tomography (MDCT) images in a sample of the Egyptian population.

**Methods:** This clinical study was conducted on 150 Egyptian subjects underwent cranium MDCT with age ranged from 21 to 60 years. The measurements used were maximum cranial breadth, minimum frontal breadth, upper facial breadth, bizygomatic breadth, orbital height, orbital breadth, parietal chord, bimastrale, maximum cranial length, basion-bregma height, cranial base length, and basion-prosthion length.

**Results:** The results revealed that stature and craniofacial measurements of males were significantly higher than those of females, all measurements were significantly positively correlated with stature in pooled cases, but the correlation coefficient differs in separate sex. Simple linear regression for stature estimation showed that the least standard error of estimate (SEE) values for the regression equations obtained when using bizygomatic breadth in pooled cases (7.9 cm) and in males only (5.7 cm), while in females using parietal chord had the least SEE (6.8 cm). Using multiple and stepwise regression analysis reported lower SEE values than simple linear regression analysis.

**Conclusion:** On the basis of this study, it is concluded that the cranial measurements obtained from MDCT images have limited utility in stature estimation among Egyptians, but could be used as an alternative method in cases where the best predictors, such as long bones, are not available.

**Keywords:** Stature estimation, Regression equations, Skull, Multi-detector computed tomography, Forensic anthropology

## Background

Forensic investigation is mainly focused on the personal identification of deceased persons, as the identity is critical in legal investigations. To identify an individual, it is necessary to establish a biological profile via the estimation of age, sex, ancestry, and stature, also known as the “big four” parameters of forensic anthropology (Divakar et al. 2015), but the

task of identification is challenging when decomposed, dismembered, or skeletal remains are recovered (Kanchan et al. 2008), stature estimation is of tremendous interest to anthropologists as it narrows down the forensic investigating process (Kumar et al. 2013).

There are two main methods available for adult stature estimation: “anatomical” and “mathematical” (Raxter et al. 2006). The anatomical method involves the direct reconstruction of stature by measuring and adding together the lengths or heights of a series of contiguous skeletal elements from the skull through the foot; it is considered most applicable and provides

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the best approximation of living stature, while the mathematical method uses regression formulae based on the correlation of individual skeletal elements to living stature (Raxter et al. 2006). Numerous studies have concluded that regression equations derived from intact long bones are accurate as long bones are highly correlated with stature. However, long bones may be unavailable, particularly in cases involving mass disasters or skeletal remains, so it is necessary to investigate the correlations between stature and the other bones (Duyar and Pelin 2003). Mathematical stature estimation is a variable among different populations and body proportions (Krishan et al. 2012). Therefore, different formulae are required for different populations and different body parts or bones (Dayal et al. 2008).

The utilization of radiological imaging especially multi-detector computed tomography (MDCT) imaging in anthropometric studies developed great attention as forensic experts can assess bones from these CT images without removing tissue, even when the subject is not skeletonized leading to a decrease in costs and time required for investigations (Franklin et al. 2016).

In cases where only the head and face are available for analysis, cranium would be of great value for person stature estimation; moreover, craniofacial structures are composed largely of hard tissues that are relatively indestructible and remain for a long period of time (Torimitsu et al. 2016).

However, little is known about the correlation between cranial measurements and stature in the Egyptian population so stature estimation formulae from the cranium are required for Egyptians. The objective of this study is to develop an anthropometric method for stature assessment in a sample of the Egyptian population by using cranium (MDCT) measurements and to derive regression equations for stature estimation.

To the best of our knowledge, it is the first study that developed regression equations for stature estimation in Egyptians from skull measurements.

## Materials and methods

This clinical study was conducted at a radiology department in Minia University Hospital after approval of the ethical committee of Minia University and obtaining written informed consent. Data were reviewed for 150 cases (80 males and 70 females) with age ranged from 21 to 60 years. All cases were subjected to head study using MDCT from December 2016 to January 2018. Subjects with skull fractures, head thermal injuries, cranial surgeries, and congenital or acquired anomalies in the craniofacial

region were excluded from the study and also when the acquired landmark was obliterated or unclear. Cases with age less than 21 years old were excluded due to marked cessation of growth and complete fusion of long bone epiphysis to diaphysis (Scheuer and Black 2004) and persons more than 60 years old as an age-related decrease in stature becomes apparent (Cline et al. 1989).

## Methods

The stature of cases was measured in centimeters using anthropometry by making the subject stand straight on a horizontal resting plane barefooted with the head in the Frankfort plane, buttocks, and heels pressed against the upright position of the instrument and the palms of the hands turned inwards and the fingers pointing downwards. Then the movable piece was brought to the vertex in the mid-sagittal plane (Numan et al. 2013).

Cranial MDCT examinations were performed using PHILIPS 16 multi-slice CT (Ingenuity Flex<sup>TM</sup>, Health care, Nederland, BV), examination protocol using 64 × 0.5-mm section collimation scanner with a gantry rotation speed of 400 ms/rotation, range of box 450–500, section thickness 0.80 mm, section time 1.60 s, standard pitch factor of 0.641, reconstruction interval 0.5 mm, and total exposure time 6.949 s. Each scan was obtained with a tube voltage of 120 kV and 399 mAs. Axial cuts were taken on the brain and skull and then reconstructed coronal and sagittal images were generated as well as 3D reformatted images that were used to obtain the acquired measures according to Moore-Jansen et al. (1994) except bimastoidale was described by Howells (1989); the used measurements landmarks were illustrated in Table 1.

Measurements obtained from three-dimensional (3D) coronal plane images are the following:

(A) Maximum cranial breadth (XCB): Distance between euryons (Fig. 1).

(B) Minimum frontal breadth (WFB): Distance between both frontotemporales (f) in the coronal plane (Fig. 1).

(C) Upper facial breadth (fmt-fmt): Distance between the two external points on the frontomalar suture (Fig. 1).

(D) Bizygomatic breadth (BZB): Distance between most lateral points on the zygomatic arches (Fig. 1).

(E) Orbital height (OBH): Distance between the superior and inferior orbital wall (Fig. 2).

(F) Orbital breadth (OBB): Distance from dacryon (d) to ectoconchion (ec); the greatest width of orbit. The diameter should cut the orbital cavity in almost two equal halves. The left orbit is used for all cases for standardization (Fig. 2).

**Table 1** Measurements landmarks

Landmark	Description
Euryon (e)	The two points on the opposite sides of the cranium that form termini of the lines of greatest cranial breadth
Frontotemporale (f)	The point where the temporal line reaches its most antero-medial position on the frontal bone
Dacryon (d)	The point of junction of the maxillary bone, lacrimal bone, and frontal bone on the medial border of the orbit
Ectoconchion (ec)	The most anterior surface of the lateral border of the orbit
Bregma (b)	The intersection of the coronal and sagittal sutures in the midline
Lambda (l)	The intersection of the sagittal and lambdoidal sutures in the midline
Glabella (g)	The most forward projecting point in the midline of the forehead at the level of the supra-orbital ridges and above the naso-frontal suture
Opisthocranium (op)	The most posterior point on the cranium
Basion (ba)	The midpoint of the anterior margin of the foramen magnum
Nasion (n)	The intersection of the internasal suture with the nasofrontal suture in the midsagittal plane
Prosthion (pr)	It is the most forward projecting point of the anterior surface of the upper jaw, in the mid-sagittal plane

(J) Parietal chord (PAC): Distance from bregma (b) to lambda (l) (Fig. 3).

(H) Bimastoidale (BMS): Distance between points of the mastoid processes (Fig. 4).

The measurement obtained from lateral 3D images is as follows:

(I) Maximum cranial length (GOL): Distance between glabella (g) and opisthocranium (Fig. 5).

Measurements obtained from the mid-sagittal plane are the following (Fig. 6):

(A) Basion-bregma height (BBH): Distance from basion to bregma

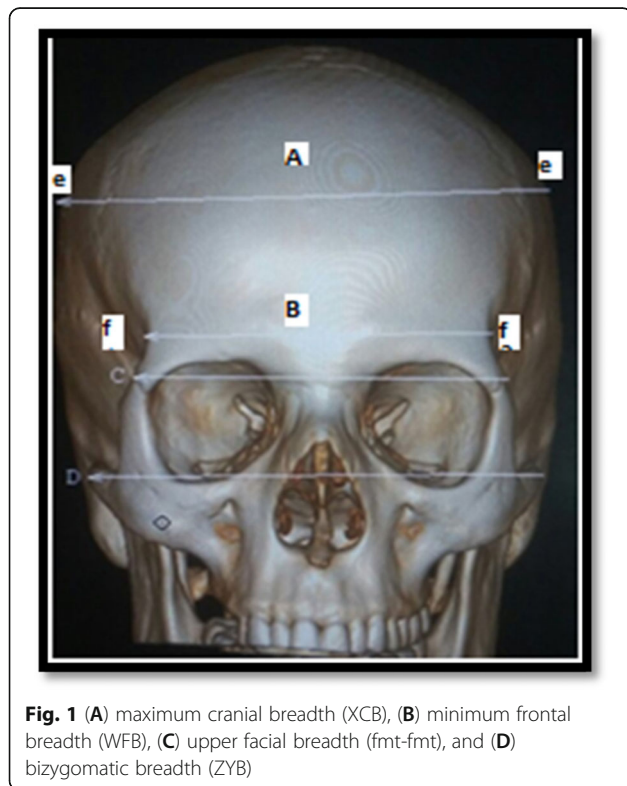
(B) Cranial base length (BNL): Distance from basion to nasion

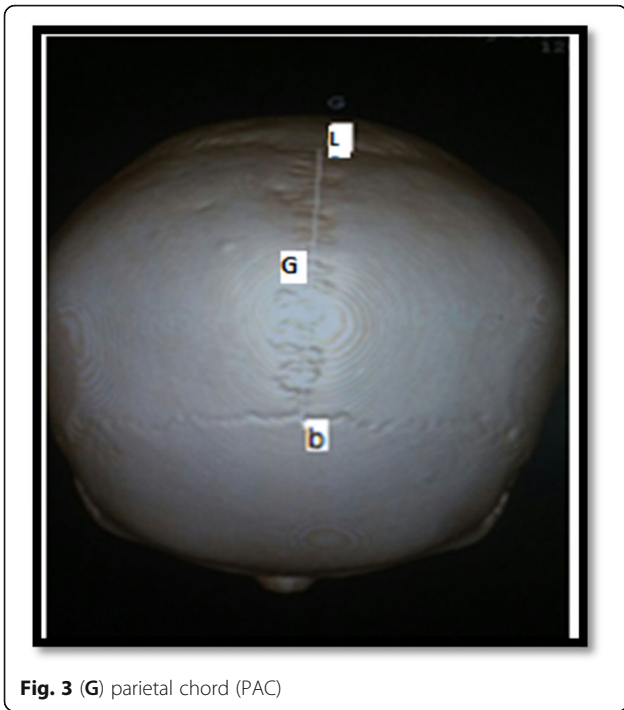
(C) Basion-prosthion length (BPL): Distance from basion to prosthion

The measurements of 10 CT images from randomly selected subjects were performed repeatedly by both the first and fourth co-author to assess inter- and intra-observer error.

**Statistical analysis**

The collected data were coded, tabulated, and statistically analyzed using SPSS program (Statistical



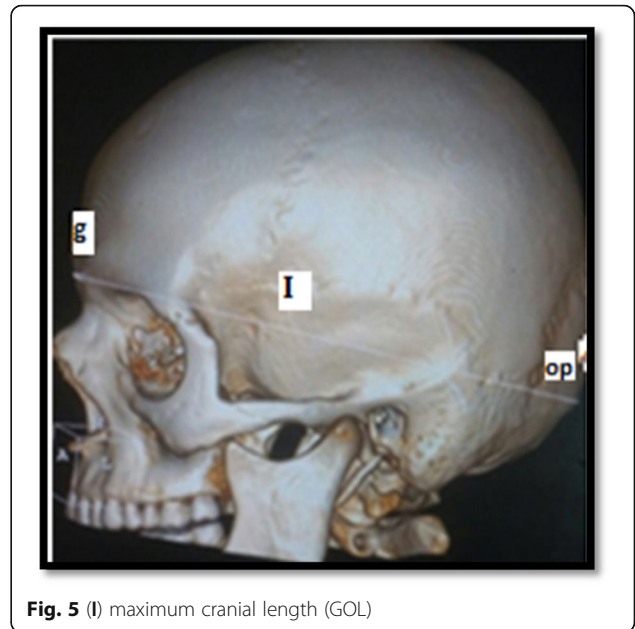


**Fig. 3 (G)** parietal chord (PAC)

Package for Social Sciences) software version 24. Descriptive statistics were done for numerical data by mean, standard deviation, and minimum and maximum of the range. Analyses were done for parametric quantitative variables between the two groups using independent samples *t* test. To assess intra- and inter-observer errors, the relative technical error of measurement (rTEM %) and the coefficient of

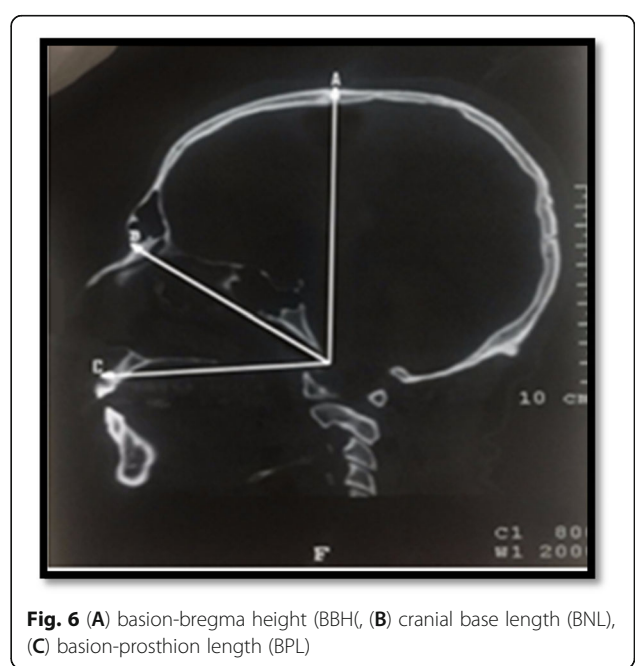


**Fig. 4 (H)** bimastoidale (BMS)



**Fig. 5 (I)** maximum cranial length (GOL)

reliability (*R*) were calculated. The acceptance ranges of rTEM in a beginner anthropometrist for intra- and inter-observe errors were < 1.5% and < 2.0%, respectively. An *R* value, a proportion of the between-subject variance that is free of measurement errors, of > 0.75 was considered sufficiently precise (Jamaiyah et al. 2010). Correlation between two quantitative variables was done by using Pearson's correlation coefficient. Simple, multiple, and multiple stepwise linear regression analyses were used to determine the equations that predict the person's



**Fig. 6 (A)** basion-bregma height (BBH), **(B)** cranial base length (BNL), **(C)** basion-prosthion length (BPL)



stature using different skull measures, standard error of estimation (SEE, centimeter) was calculated for each formula to evaluate the significance of the regression equations. The adjusted coefficient of determination ( $R^2_{adj}$ ) was calculated for multiple and multiple stepwise regression analysis, while the coefficient of determination ( $R^2$ ) was calculated for simple regression analysis formulae.

**Results**

The results revealed that the mean age of tested cases was  $38.7 \pm 12.3$  years. The stature of cases ranged from 143 to 184 cm with a mean of  $163.8 \pm 9.5$ . The descriptive statistics of different skull measurements among males and females were shown in Table 2, Independent samples *T* test revealed a significant increase in all skull measurements in males in comparison with females (Table 2). Intra-observer precisions of skull measurements were  $rTEM < 1.5\%$  (0.61–1.07%) and  $R > 0.95$  (0.981–0.994); inter-observer precisions were  $rTEM < 2.0\%$  (0.67–1.76%) and  $R > 0.95$  (0.957–0.986) as shown in Table 3.

The correlation coefficient between the stature and all skull measurements in pooled cases reported that all skull measurements had significant positive correlation, while in males, ZYB, GOL XCB, BPL, and OBH measurements showed significant positive correlation with stature, and in females, XCB, ZYB, OBB, PAC, and GOL measurements showed

**Table 3** Intra-observer and inter-observer of relative technical error of measurements (rTEM %) and coefficient of reliability (R) for all skull measurements (n = 10 cases)

Measurements	Intra-observer		Inter-observer	
	rTEM (%)	R	rTEM (%)	R
XCB	0.61	0.981	0.67	0.986
WFB	0.62	0.984	1.01	0.957
Fmt-fmt	1.07	0.981	1.42	0.963
ZYB	0.92	0.984	1.76	0.962
OBH	1	0.991	1.72	0.982
OBB	0.88	0.987	1.34	0.957
PAC	0.99	0.992	1.13	0.980
BMS	1.17	0.985	1.92	0.942
GOL	0.82	0.982	0.91	0.983
BBH	0.7	0.990	1.06	0.970
BNL	0.88	0.987	1.33	0.974
BPL	0.71	0.994	1.23	0.979

XCB maximum cranial breadth, WFB minimum frontal breadth, *fmt-fmt* upper facial breadth, ZYB bizygomatic breadth, OBH orbital height, OBB orbital breadth, PAC parietal chord, BMS bi-mastoidale, GOL maximum cranial length, BBH basion-bregma, BNL cranial base length, BPL basion-prosthion length

significant positive correlation with stature as shown in Table 4.

Tables 5, 6, and 7 showed the results of the simple linear regression equations for stature estimation derived for each of the cranial measurements in pooled cases, males only and females only, respectively. The

**Table 2** Descriptive statistics and independent sample *t* test of stature and skull measurements in pooled cases (n = 150 cases), males (n = 80 cases), and females (n = 70 cases)

Measurements	Mean ± SD		Minimum		Maximum		P value
	Male	Female	Male	Female	Male	Female	
Age	38.3 ± 12	39.2 ± 12.8	21 y	21 y	60 y	60 y	< 0.001*
Stature	169.9 ± 6.4	156.9 ± 7.5	153 cm	143 cm	184 cm	175 cm	< 0.001*
XCB	14.4 ± 1.4	13.3 ± 1.1	11 cm	9.5 cm	18.9 cm	17.2 cm	< 0.001*
WFB	9.9 ± 0.5	9.4 ± 0.4	8.6 cm	8.6 cm	11.1 cm	11 cm	< 0.001*
Fmt-fmt	10.9 ± 0.6	10.4 ± 0.7	9.7 cm	8.7 cm	12.8 cm	13.3 cm	< 0.001*
ZYB	14.6 ± 2.1	13 ± 1.4	11.2 cm	11.3 cm	19.4 cm	18.9 cm	< 0.001*
OBH	3.9 ± 0.3	3.7 ± 0.3	3.3 cm	3.1 cm	4.7 cm	4.8 cm	< 0.001*
OBB	4.1 ± 0.4	3.9 ± 0.3	3.4 cm	3.3 cm	4.9 cm	4.8 cm	< 0.001*
PAC	12 ± 1	11.6 ± 0.8	9.5 cm	9.9 cm	16.7 cm	14.3 cm	0.007*
BMS	10.5 ± 0.6	10 ± 0.5	8.7 cm	8.5 cm	12.4 cm	10.9 cm	< 0.001*
GOL	18.5 ± 1	17.8 ± 0.8	14.8 cm	16.3 cm	20.9 cm	20.6 cm	< 0.001*
BBH	13.9 ± 0.7	13.4 ± 0.5	11 cm	11.8 cm	15.1 cm	14.8 cm	< 0.001*
BNL	10.4 ± 1.2	10 ± 0.6	8 cm	8.6 cm	16.6 cm	11.4 cm	0.005*
BPL	10.2 ± 0.9	9.7 ± 0.5	8.3 cm	8.2 cm	15 cm	10.8 cm	< 0.001*

All skull parameters and stature were measured in centimeter

SD standard deviation, XCB maximum cranial breadth, WFB minimum frontal breadth, *fmt-fmt* upper facial breadth, ZYB bizygomatic breadth, OBH orbital height, OBB orbital breadth, PAC parietal chord, BMS bi-mastoidale, GOL maximum cranial length, BBH basion-bregma height, BNL cranial base length, BPL basion-prosthion length

**Table 4** Correlation between the stature and skull measurements in pooled cases ( $n = 150$ ), males ( $n = 80$ ), and females ( $n = 70$ ) by Pearson's correlation coefficient

	XCB	WFB	Fmt-fmt	ZYB	OBH	OBB	PAC	BMS	GOL	BBH	BNL	BPL
Pooled cases												
R	0.519	0.384	0.345	0.555	0.382	0.394	0.268	0.373	0.497	0.348	0.221	0.376
P value	< 0.001*	< 0.001*	< 0.001*	< 0.001*	< 0.001*	< 0.001*	< 0.001*	< 0.001*	< 0.001*	< 0.001*	0.006*	< 0.001*
In males												
R	0.336	0.091	0.068	0.452	0.381	0.157	0.034	0.091	0.401	0.177	0.116	0.280
P value	0.002*	0.420	0.552	< 0.001*	< 0.001*	0.165	0.765	0.420	< 0.001*	0.115	0.306	0.012*
In female												
R	0.395	0.120	0.190	0.387	0.154	0.362	0.405	0.173	0.308	0.114	0.093	0.158
P value	0.001*	0.321	0.115	0.001*	0.204	0.002*	0.001*	0.151	0.010*	0.347	0.442	0.191

*r*, correlation coefficient; weak ( $R = 0.20-0.39$ ), moderate ( $R = 0.40-0.59$ ), strong ( $R = 0.60-0.79$ ), very strong  $> 0.8$

XCB maximum cranial breadth, WFB minimum frontal breadth, *fmt-fmt* upper facial breadth, ZYB bizygomatic breadth, OBH orbital height, OBB orbital breadth, PAC parietal chord, BMS bi-mastoidale, GOL maximum cranial length, BBH basion-bregma, BNL cranial base length, BPL basion-prosthion length

\*Significant difference,  $P$  value < 0.05

**Table 5** Simple linear regression analysis to predict stature from skull measurements in pooled cases ( $n = 150$ )

Measures	Unstandardized coefficients		P value	$R^2$	SEE (cm)	Regression equation
	B	Std. error				
XCB	3.57	0.48	< 0.001*	0.269	8.12	$S = 114.4 + (3.57 \times XCB)$
Constant	114.4	6.73	< 0.001*			
WFB	6.95	1.38	< 0.001*	0.147	8.77	$S = 96.9 + (6.95 \times WFB)$
Constant	96.9	13.27	< 0.001*			
Fmt-fmt	4.57	1.02	< 0.001*	0.119	8.92	$S = 115.2 + (4.57 \times \text{fmt-fmt})$
Constant	115.2	10.91	< 0.001*			
ZYB	2.7	0.33	< 0.001*	0.308	7.9	$S = 126.5 + (2.7 \times ZYP)$
Constant	126.5	4.64	< 0.001*			
OBH	11.18	2.22	< 0.001*	0.146	8.78	$S = 121.1 + (11.18 \times OBH)$
Constant	121.1	8.51	< 0.001*			
OBB	11.07	2.12	< 0.001*	0.155	8.73	$S = 119.6 + (11.07 \times OBB)$
Constant	119.6	8.5	< 0.001*			
PAC	2.8	0.83	0.001*	0.072	9.15	$S = 130.7 + (2.8 \times PAC)$
Constant	130.7	9.85	< 0.001*			
BMS	5.84	1.2	< 0.001*	0.139	8.81	$S = 104 + (5.84 \times BM)$
Constant	104	12.28	< 0.001*			
GOL	4.85	0.7	< 0.001*	0.247	8.24	$S = 75.7 + (4.85 \times GOL)$
Constant	75.7	12.67	< 0.001*			
BBH	4.83	1.07	< 0.001*	0.121	8.91	$S = 98 + (4.83 \times BBH)$
Constant	98	14.61	< 0.001*			
BNL	2.05	0.74	0.006*	0.049	9.26	$S = 142.9 + (2.05 \times BNL)$
Constant	142.9	7.62	< 0.001*			
BPL	4.54	0.92	< 0.001*	0.142	8.8	$S = 118.5 + (4.54 \times BPL)$
Constant	118.5	9.2	< 0.001*			

$R^2$ , co-efficient determination; < 0.1 (trivial), 0.1-0.3 (small to medium), 0.3-0.5 (medium to large), > 0.5 (large to very large)

SEE standard error of estimate

\*Significant level at  $P$  value < 0.05

**Table 6** Simple linear regression analysis to predict stature from skull measurements in males (N = 80)

Skull measurements	Unstandardized coefficients		P value	R <sup>2</sup>	SEE (cm)	Regression equation
	B	Std. error				
XCB	1.54	0.49	< 0.001*	0.113	6.03	S = [147.8 + (1.54 × XCB)]
Constant	147.8	7.07	< 0.001*			
ZYB	1.4	0.31	< 0.001*	0.204	5.71	S = [149.5 + (1.4 × ZYB)]
Constant	149.5	4.6	< 0.001*			
OBH	7.11	1.95	< 0.001*	0.145	5.92	S = [142.2 + (7.11 × OBH)]
Constant	142.2	7.65	< 0.001*			
GOL	2.65	0.69	< 0.001*	0.160	5.87	S = [120.9 + (2.65 × GOL)]
Constant	120.9	12.72	< 0.001*			
BPL	2	0.78	0.012*	0.078	6.15	S = [149.5 + (2 × BPL)] ± 6.15
Constant	149.5	7.97	< 0.001*			

R<sup>2</sup>, co-efficient determination; < 0.1 (trivial), 0.1–0.3 (small to medium), 0.3–0.5 (medium to large), > 0.5 (large to very large)

SEE standard error of estimate

\*Significant level at P value < 0.05

most predictable measurement was ZYB in pooled cases and in males only with the lowest SEE 7.9 cm and 5.71 cm, respectively, while PAC was the most predictable and had the least SEE (6.86 cm) in females.

The results of multiple linear regression equations were shown in Table 8 which showed less SEE values than simple linear regression equations in pooled cases, males and females.

Tables 9, 10, and 11 showed results of stepwise multiple linear regression analysis in pooled cases, males and females, respectively, which demonstrated nearly the same SEE values of multiple linear regression with less measurements used that is valuable especially in fragmented skull.

## Discussion

Estimation of stature is one of the essential components in establishing identity in forensic anthropology as well as cranium is an important bone that has generated great interest among forensic anthropologists; furthermore, MDCT has generated valuable rule in forensic anthropology. So, this study aimed to stature estimation using cranial measurements in MDCT images.

The mean stature in this study was more in males (169.9 ± 6.4 cm) in comparison with females (156.9 ± 7.5 cm) which agrees with other studies done on Sudanese by Ahmed and Taha (2016), Indian by Shrestha et al. (2015), and Japanese population by Torimitsu et al. (2016). However, the mean values of

**Table 7** Simple linear regression analysis to predict stature from skull measurements in females (n = 70)

Skull measurements	Unstandardized coefficients		P value	R <sup>2</sup>	SEE (cm)	Regression equation
	B	Std. error				
XCB	2.7	0.76	0.001*	0.156	6.9	S = 121.1 + (2.7 × XCB)
Constant	121.1	10.13	< 0.001*			
ZYB	2.06	0.6	0.001*	0.150	6.92	S = 130.2 + (2.06 × ZYP)
Constant	130.2	7.77	< 0.001*			
OBB	9.97	3.11	0.002*	0.131	7	S = 118.3 + (9.97 × OBB)
Constant	118.3	12.08	< 0.001*			
PAC	3.74	1.03	0.001*	0.164	6.86	S = 113.3 + (3.74 × PAC)
Constant	113.3	11.96	< 0.001*			
GOL	2.77	1.04	0.010*	0.095	7.14	S = 107.6 + (2.77 × GOL)
Constant	107.6	18.5	< 0.001*			

R<sup>2</sup>, co-efficient determination; < 0.1 (trivial), 0.1–0.3 (small to medium), 0.3–0.5 (medium to large), > 0.5 (large to very large)

SEE standard error of estimate

\*Significant level at P value < 0.05

**Table 8** multiple linear regression analysis to predict stature from skull measurements in pooled cases ( $n = 150$ ), males ( $n = 80$ ), and females ( $n = 70$ )

Subjects	Regression equations	$R^2_{adj}$	SEE (cm)
Pooled cases	$S = [3.34 + (1.32 \times XCB) + (1.52 \times WFB) + (0.05 \times \text{fmt-fmt}) + (1.49 \times ZYB) + (5.54 \times OBH) + (3.23 \times OBB) + (0.89 \times PAC) + (1.76 \times BMS) + (0.49 \times GOL) + (0.91 \times BBH) + (-0.01 \times BNL) + (2.25 \times BPL)]$	0.566	6.21
Males	$S = [10.28 + (0.87 \times XCB) + (1.5 \times ZYB) + (5.52 \times OBH) + (0.85 \times GOL) + (2.26 \times BPL)]$	0.455	4.7
Females	$S = [71.9 + (1.67 \times XCB) + (0.75 \times ZYB) + (7.37 \times OBB) + (1.93 \times PAC) + (0.12 \times GOL)]$	0.285	6.3

$R^2_{adj}$  adjusted co-efficient determination; < 0.1 (trivial), 0.1–0.3 (small to medium), 0.3–0.5 (medium to large), > 0.5 (large to very large)

SEE standard error of estimate

\*Significant level at  $P$  value < 0.05

Egyptian stature were lower than those of Sudanese and slightly higher than that of Indian and Japanese populations. Also, the means of skull measurements were significantly higher in males than in females; this is supported by most studies done in different populations.

The rTEMs and  $R$  values indicate that a technical error was likely to be accepted in this study which suggested that bone measurements using MDCT images can be performed without significant technical error; this coincides with Torimitsu et al. (2016).

The present findings indicate that stature and all skull measurements used are positively and significantly correlated in all Egyptian cases, but these results differed by sex. In males, XCB, ZYB, OBH, GOL, and BPL measurements showed significant correlation with stature and in females, XCB, ZYB, OBB, PAC, and GOL measurements had significant correlation with stature; the higher correlation coefficients were observed for pooled cases than for males or females may be attributed to the larger sample size (the sum of males and females). Acharya et al. (2017) studied 113 dry skulls of South Indian males and concluded that GOL had significant positive strong correlation and XCB had moderate correlation, while fmt-fmt and PAC had a fair correlation with stature. Torimitsu et al. (2016) who reviewed 228 MDCT images of the Japanese population found that there were significant positive correlations between stature and all the five measurements used in all subjects, among males, all cranial measurements except PAC were significantly correlated with stature. However, among females, only the ZYB and XCB measurements were significantly correlated with stature which disagrees with the present results. However, in Colombians, González-Colmenares et al. (2016) reported that the correlation coefficients were found to be low in males. Among females, the co-efficient of correlation was not found to be statistically significant for stature and cranial measurements; this may be attributed to

the a decrease number of Colombian female samples used (16 dry skulls) which differ from these results.

Studies among Northwest Indian (Sahni et al. 2010) and Indo-Mauritian (Agnihotri et al. 2011) populations showed low correlations between stature and seven facial measurements ( $r < 0.270$  for males and  $r < 0.195$  for females) and 14 cephalo-facial measurements ( $r < 0.494$  for males and  $< 0.382$  for females), respectively. Therefore, the different cephalo-facial measurements used in different studies are not similar with regard to stature; they differ between the sexes and among different populations as genetic factors, environmental and nutritional factors affect the growth of the skull (Duren et al. 2013).

Regarding SEE of simple linear regression analysis reported in this study ranged from 7.9 cm to 9.26 cm, 5.71 cm to 6.15 cm, and 6.86 cm to 7.14 cm in pooled cases, males and females, consequently, The high SEE values noted for combined sex were probably due to increased variance in the sample, and these results were comparable with Torimitsu et al. (2016) who reported SEE values in Japanese ranged from 6.918 cm to 9.123 cm in pooled cases, 5.581 cm to 5.956 cm in males, and 6.451 cm to 7.142 cm in females. Also, Chiba and Terazawa (1998) who investigated Japanese dry skull at autopsy reported SEE values ranged from 7.95 cm to 8.59 in pooled cases and 6.96 cm to 7.12 cm and 6.71 cm to 6.97 cm in males and females consequently which coincides with these results. Shreshta et al. (2015) studied 200 skulls of Nepalese cadavers and reported a minimum SEE value with GOL, 7.437 cm in pooled cases, and 6.894 cm in males and in females; the least SEE was 5.692 cm.

In Sudanese, Ahmed and Taha (2016) reported lower SEE values ranged from 5.745 cm to 6.028 cm and from 5.587 cm to 5.903 cm in males and in females, consequently.

Multiple studies investigated a sample of Indian male skulls percutaneously as Krishan (2008) and



**Table 9** Multiple stepwise linear regression analysis of skull measurements for stature estimation in pooled cases ( $n = 150$ )

Model	Measures	Unstandardized Coefficients		P value	$R^2_{adj}$	SEE	Regression equation
		B	Std. error				
1	ZYB	2.7	0.33	< 0.001*	0.304	7.9	$S = 126.5 + (2.7 \times ZYB)$
	Constant	126.5	4.64	< 0.001*			
2	ZYB	2.03	0.33	< 0.001*	0.409	7.3	$S = 101.8 + (2.03 \times ZYB) + (2.46 \times XCB)$
	XCB	2.46	0.47	< 0.001*			
	Constant	101.8	6.38	< 0.001*			
3	ZYB	1.88	0.31	< 0.001*	0.484	6.8	$S = 72.1 + (1.88 \times ZYB) + (2.31 \times XCB) + (3.39 \times BPL)$
	XCB	2.31	0.44	< 0.001*			
	BPL	3.39	0.72	< 0.001*			
	Constant	72.1	8.66	< 0.001*			
4	ZYB	1.82	0.3	< 0.001*	0.428	6.5	$S = 56.3 + (1.82 \times ZYB) + (2.04 \times XCB) + (2.88 \times BPL) + (6.34 \times OBB)$
	XCB	2.04	0.43	< 0.001*			
	BPL	2.88	0.7	< 0.001*			
	OBB	6.34	1.65	< 0.001*			
	Constant	56.3	9.25	< 0.001*			
5	ZYB	1.8	0.29	< 0.001*	0.547	6.4	$S = 47.5 + (1.8 \times ZYB) + (1.78 \times XCB) + (2.79 \times BPL) + (5.25 \times OBB) + (4.69 \times OBH)$
	XCB	1.78	0.43	< 0.001*			
	BPL	2.79	0.69	< 0.001*			
	OBB	5.25	1.67	0.002*			
	OBH	4.69	1.75	0.008*			
	Constant	47.5	9.64	< 0.001*			
6	ZYB	1.69	0.29	< 0.001*	0.562	6.3	$S = 31.5 + (1.69 \times ZYB) + (1.62 \times XCB) + (2.5 \times BPL) + (4.27 \times OBB) + (5.51 \times OBH) + (2.29 \times BMS)$
	XCB	1.62	0.43	< 0.001*			
	BPL	2.5	0.69	< 0.001*			
	OBB	4.27	1.69	0.013*			
	OBH	5.51	1.75	0.002*			
	BMS	2.29	0.94	0.016*			
	Constant	31.5	11.54	0.007*			

$R^2_{adj}$ , adjusted co-efficient determination; < 0.1 (trivial), 0.1–0.3 (small to medium), 0.3–0.5 (medium to large), > 0.5 (large to very large)

SEE standard error of estimate

\*Significant level at  $P$  value < 0.05

Krishan and Kumar (2007) and revealed SEE values 3.726 cm to 5.820 cm, 4.41 cm to 7.21 cm, consequently. Also, Patil and Mody (2005) used a lateral cephalogram of 150 adults from Central India and indicated that GOL can be used to estimate the stature; however, the SEE and  $R$  values for their regression equations were not provided for comparison. The SEE values of Sahni et al. (2010) who studied facial measurements of 300 adults in Northwest India were much less than the values reported in this study, 3.56 cm to 3.70 cm in males and 2.90 to 2.95 in females.

The results of SEE of multiple linear regression analysis and stepwise regression analysis showed lower SEE values than simple linear regression

analysis, 6.3 cm in pooled cases, 4.7 cm in males, and 6.3 cm in females; this goes in line with most studies mentioned above which concluded the same as illustrated in Table 12.

Genetics and environment may affect skeletal development; thus, when cephalometric measurements are used to determine sex or stature, it is necessary to develop specific formulae for each population (Vercellotti et al. 2009). The great result variability may be due to the difference in the measurements used in each study and techniques used for obtaining these measures in either MDCT, percutaneous, or dry bones. Although Franklin et al. (2013) stated that the precision of MDCT-derived measurements is high enough for the error to be negligible

**Table 10** Multiple stepwise linear regression analysis of skull measurements for stature estimation in males (n = 80)

Model	Measures in males	Unstandardized coefficients		P value	R <sup>2</sup> <sub>adj</sub>	SEE	Regression equation
		B	Std. error				
1	ZYB	1.4	0.31	< 0.001*	0.194	5.71	S = [149.5 + (1.4 × ZYP)]
	Constant	149.5	4.6	< 0.001*			
2	ZYB	1.35	0.29	< 0.001*	0.318	5.26	S = [123.8 + (1.35 × ZYP) + (6.77 × OBH)]
	OBH	6.77	1.73	< 0.001*			
	Constant	123.8	7.84	< 0.001*			
3	ZYB	1.43	0.27	< 0.001*	0.399	4.93	S = [102.5 + (1.43xZYP) + (6.42xOBH) + (2.11xBPL)]
	OBH	6.42	1.63	< 0.001*			
	BPL	2.11	0.63	0.001*			
	Constant	102.5	9.7	< 0.001*			
4	ZYB	1.3	0.27	< 0.001*	0.440	4.76	S = [82.5 + (1.3 × ZYP) + (5.87 × OBH) + (1.74 × BPL) + (1.5 × GOL)]
	OBH	5.87	1.59	< 0.001*			
	BPL	1.74	0.62	0.006*			
	GOL	1.5	0.59	0.013*			
	Constant	82.5	12.2	< 0.001*			

R<sup>2</sup><sub>adj</sub>, adjusted co-efficient determination; < 0.1 (trivial), 0.1–0.3 (small to medium), 0.3–0.5 (medium to large), > 0.5 (large to very large)

SEE standard error of estimate

\*Significant level at P value < 0.05

compared to the sample variance but also found that traditional bone measurements are more precise than their MDCT counterparts, so this point must be further investigated.

**Conclusion**

On the basis of simple regression equations developed from this study, it is concluded that bizygomatic breadth could be used for stature estimation in pooled cases and in males, while parietal chord in females and using multiple and multiple stepwise

regression equations had lower SEE values. However R<sup>2</sup> values are too low which makes cranial MDCT images have limited utility; hence, these should be used a last resort and alternative method for stature estimation in Egyptians when more accurate skeletal elements such as intact long bones are not available for investigations.

The age groups used (21–60 years) are considered a limitation in this study as age groups less than the age of epiphyseal closure must be investigated to assess the reliability of using cranial MDCT for their

**Table 11** Multiple stepwise linear regression analysis of skull measurements for stature estimation in females (n = 70)

Model	Measures in females	Unstandardized Coefficients		P value	R <sup>2</sup> <sub>adj</sub>	SEE	Regression equation
		B	Std. error				
1	PAC	3.74	1.03	0.001*	0.152	6.86	S = [113.3 + (3.74 × PAC)]
	Constant	113.3	11.96	< 0.001*			
2	PAC	3.26	0.99	0.002*	0.229	6.54	S = [87 + (3.26 × PAC) + (8.27 × OBB)]
	OBB	8.27	2.96	0.007*			
	Constant	87	14.79	< 0.001*			
3	PAC	2.3	1.02	< 0.001*	0.290	6.29	S = [72.59 + (2.3 × PAC) + (8.12 × OBB) + (1.94 × XCB)]
	OBB	8.12	2.84	0.027*			
	XCB	1.94	0.75	0.005*			
	Constant	72.59	15.23	< 0.001*			

R<sup>2</sup><sub>adj</sub>, adjusted co-efficient determination; < 0.1(trivial), 0.1–0.3 (small to medium), 0.3–0.5 (medium to large), > 0.5 (large to very large)

SEE standard error of estimate

\*Significant difference at P value < 0.05

**Table 12** demonstrating SEE in different studies among different population

Study	Population	Method	Sex	SEE (simple) (cm)	SEE (multiple) (cm)
Krishan and Kumar (2007)	North Indian	Percutaneous	Male	4.41	–
Krishan (2008)	North Indian	Percutaneous	Male	4.136	–
Sahni et al. (2010)	Northwest Indian	Percutaneous		3.56	3.569
			Male	2.90	2.880
Acharya et al. (2017)	South India	Direct bone	Female	4.90	4.45
Shrestha et al. (2015)	Nepalese	Direct bone	Male	6.894	6.640
			Female	5.692	5.538
			Pooled	7.437	6.943
Chiba and Terazawa (1998)	Japanese	Direct bone	Male	7.09	6.96
			Female	6.59	6.71
			Pooled	8.14	7.95
Torimitsu et al. (2016)	Japanese	3D CT bone	Male	5.581	5.406
			Female	6.451	6.314
			Pooled	6.918	6.394
Ahmed and Taha (2016)	Sudanese	Percutaneous	Male	5.698	5.532
			Female	5.587	5.253
			Pooled	–	–
The current study	Egyptian	3D CT bone	Male	5.71	4.76
			Female	6.86	6.3
			Pooled	7.9	6.3

stature estimation, and also, studying the effect of age on stature and skull measurements must be further investigated. Software programs for landmark location determination are not available. Another limitation that the regression equations developed from this study is untested for its validation; this point needs future consideration.

#### Abbreviations

BBH: Basion-bregma height; BMS: Bi-mastoidale; BNL: Cranial base length; BPL: Basion-prosthion length; fnt-fmt: Upper facial breadth; GOL: Maximum cranial length; MDCT: Multi-detector computed tomography; OBB: Orbital breadth; OBH: Orbital height; PAC: Parietal chord; SEE: Standard error of estimate; WFB: Minimum frontal breadth; XCB: Maximum cranial breadth; ZYB: Bizygomatic breadth

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HA and OH are responsible for the interpretation of results and for writing and final revision of the paper. RA and EG are responsible for obtaining the measures of the research. All authors read and approved the final manuscript.

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#### Competing interests

The authors declare that they have no competing interest.

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